

Proof of Lemma 1

Proof. For fixed $\lambda > 0$, $0 < \alpha < 1$, if $\hat{\beta}_i \neq \hat{\beta}_j$, take $\hat{\beta}^*$ such that $\hat{\beta}_k^* = \hat{\beta}_k$ when $k \neq i$ and $k \neq j$. When $k = i$ or $k = j$, let $\hat{\beta}_k^* = \frac{1}{2}(\hat{\beta}_i + \hat{\beta}_j)$. Since $\mathbf{x}_i = \mathbf{x}_j$, $\mathbf{X}\hat{\beta}^* = \mathbf{X}\hat{\beta}$. As for $1 \leq i \leq n$, $X_i\hat{\beta}^* = X_i\hat{\beta}$. While the objective function $\mathbb{L}(\lambda, \alpha, \beta)$ is strictly convex, so $\mathbb{L}(\lambda, \alpha, \hat{\beta}^*) < \mathbb{L}(\lambda, \alpha, \hat{\beta})$. At the same time, $\hat{\beta}$ satisfies equation (4), which leads to a contradiction. Consequently, $\hat{\beta}_i = \hat{\beta}_j$ must hold. If $\hat{\beta}_i \hat{\beta}_j < 0$, take the same $\hat{\beta}^*$ again. According to the triangle inequality, $|\hat{\beta}^*|_1 < |\hat{\beta}|_1$, meaning that $\hat{\beta}$ is not a solution for lasso problem, which is also a contradiction. Thus $\hat{\beta}_i \hat{\beta}_j \geq 0$ holds. Also, $X_i\hat{\beta}^* = X_i\hat{\beta}$, $1 \leq i \leq n$, for $\hat{\beta}^*$ defined in the condition. So $\hat{\beta}^*$ is also minimizer of $\mathbb{L}(\lambda, \alpha, \beta)$. ■

Proof of Theorem 1

Proof. Since $\hat{\beta}_i(\lambda, \alpha) \hat{\beta}_j(\lambda, \alpha) > 0$, both $\hat{\beta}_i$ and $\hat{\beta}_j$ are non-zeros and $\text{sign}(\hat{\beta}_i) = \text{sign}(\hat{\beta}_j)$, because $\hat{\beta} = \arg \min_{\beta} \mathbb{L}(\lambda, \alpha, \beta)$, $\hat{\beta}$ satisfies $\frac{\partial \mathbb{L}}{\partial \beta_k} \big|_{\beta=\hat{\beta}} = 0$ if $\hat{\beta}_k(\lambda, \alpha) \neq 0$. Hence,

$$\left(-y + \frac{e^{\mathbf{X}\hat{\beta}}}{1 + e^{\mathbf{X}\hat{\beta}}}\right)^T \mathbf{x}_k + \lambda \alpha \text{sign}(\hat{\beta}_k) + 2\lambda(1 - \alpha)\hat{\beta}^T L_k = 0, \quad (8)$$

where L_k is the k th column of the Laplacian matrix L . The reduction of left hand side of equation (8) is given in the next subsection. Hence

$$\left(-y + \frac{e^{\mathbf{X}\hat{\beta}}}{1 + e^{\mathbf{X}\hat{\beta}}}\right)^T \mathbf{x}_i + \lambda \alpha \text{sign}(\hat{\beta}_i) + 2\lambda(1 - \alpha)\hat{\beta}^T L_i = 0 \quad (9)$$

$$\left(-y + \frac{e^{\mathbf{X}\hat{\beta}}}{1 + e^{\mathbf{X}\hat{\beta}}}\right)^T \mathbf{x}_j + \lambda \alpha \text{sign}(\hat{\beta}_j) + 2\lambda(1 - \alpha)\hat{\beta}^T L_j = 0 \quad (10)$$

Subtracting (10) from (9), we get

$$\left(-y + \frac{e^{\mathbf{X}\hat{\beta}}}{1 + e^{\mathbf{X}\hat{\beta}}}\right)^T (\mathbf{x}_i - \mathbf{x}_j) + 2\lambda(1 - \alpha)\hat{\beta}^T (L_i - L_j) = 0 \quad (11)$$

According to property of the Laplacian matrix L ,

$$\hat{\beta}^T (L_i - L_j) = \hat{\beta}_i - \hat{\beta}_j \quad (12)$$

From (11), (12) along with Cauchy-Schwartz inequality and the property of L_1 norm, we know

$$|\hat{\beta}_i - \hat{\beta}_j| \leq \frac{1}{2\lambda(1 - \alpha)} \left| y - \frac{e^{\mathbf{X}\hat{\beta}}}{1 + e^{\mathbf{X}\hat{\beta}}} \right|_1 \cdot |\mathbf{x}_i - \mathbf{x}_j|_1 \quad (13)$$

Since $\hat{\beta}$ is the minimizer, the residual $|y - \frac{e^{\mathbf{X}\hat{\beta}}}{1 + e^{\mathbf{X}\hat{\beta}}}|_1 < |y - \frac{e^{\mathbf{X}\beta}}{1 + e^{\mathbf{X}\beta}}|_1 \leq |y|_1 + |\frac{e^{\mathbf{X}\beta}}{1 + e^{\mathbf{X}\beta}}|_1$ which tends to $|y|_1$ as β tends to negative infinity. So

$$|\hat{\beta}_i - \hat{\beta}_j| < \frac{1}{2\lambda(1 - \alpha)} |y|_1 \cdot |\mathbf{x}_i - \mathbf{x}_j|_1 \quad (14)$$

Since $\mathbf{X}$ are standardised, $|x_i - x_j|_1 = \sqrt{2(1 - \rho)}$, dividing both sides of (14) by $|y|_1$ leads to equation (5).

■

Proof of Lemma 2

Proof. Taking the partial derivative of $l(\lambda, \alpha, \beta)$ with respect to β , we get

$$\frac{\partial l}{\partial \beta} = -\mathbf{X}^T \mathbf{y} + \mathbf{X}^T \cdot \frac{e^{\mathbf{X}\beta}}{1 + e^{\mathbf{X}\beta}} + \lambda \alpha \text{sign}(\beta) + 2\lambda(1 - \alpha)L\beta. \quad (15)$$

In (15), the sign of the divisor denotes element-wise division. From KKT conditions, we know if $\beta = 0$, λ will make $\frac{\partial l}{\partial \beta}$ zero. That is,

$$\frac{\partial l}{\partial \beta}|_{\beta=0} = -\mathbf{X}^T \mathbf{y} + \frac{1}{2} \mathbf{X}^T \cdot \mathbf{1}_n + \lambda \alpha \mathbf{1}_p = \mathbf{0}_p, \quad (16)$$

where $\mathbf{0}_p$ is a column vector which contains p zero elements. Hence,

$$\lambda \alpha \mathbf{1}_p = \mathbf{X}^T \mathbf{y} - \frac{1}{2} \mathbf{X}^T \cdot \mathbf{1}_n. \quad (17)$$

The dimensions of $\mathbf{1}_p$ and $\mathbf{1}_n$ here are $p \times 1$ and $n \times 1$ respectively. Taking L_∞ norm on both sides of (17) leads to

$$\alpha \lambda = |\mathbf{X}^T \mathbf{y} - \frac{1}{2} \mathbf{X}^T \cdot \mathbf{1}_n|_\infty \leq |\mathbf{X}^T \mathbf{y}|_\infty + \frac{1}{2} |\mathbf{X}^T \cdot \mathbf{1}_n|_\infty. \quad (18)$$

Thus

$$\lambda \leq \frac{|\mathbf{X}^T \mathbf{y}|_\infty + \frac{1}{2} |\mathbf{X}^T \cdot \mathbf{1}_n|_\infty}{\alpha} \quad (19)$$

and we can take

$$\lambda_{\max} = \frac{2|\mathbf{X}^T \mathbf{y}|_\infty + |\sum_{i=1}^n X_i^T|_\infty}{2\alpha} \quad (20)$$

which is the condition in Lemma 3. ■